# Stochastic Models on Time to Recruitment in a Two Grade Manpower System using Different Policies of Recruitment 

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#### Abstract

In this paper a two grade organization in which depletion of manpower occurs due to its policy decisions is considered. Two mathematical models are constructed employing two different univariate recruitment policies, based on shock model approach. The mean and variance of the time to recruitment are obtained for both the models under different conditions. The analytical results are numerically illustrated and relevant conclusions are presented.


Key Words: Two grade system, Shock models, Univariate policies of recruitment, Mean and variance of the time to recruitment.
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## Introduction

Consider an organization having two grades in which depletion of manpower occurs at every decision epoch. In the univariate policy of recruitment based on shock model approach recruitment is made as and when the cumulative loss of manpower crosses a threshold. Employing this recruitment policy, expected time to recruitment is obtained under different conditions for several models in [1], [3], [4], [6], [8], [9], [10] and [11]. Recently in [2], Dhivya and Srinivasan have constructed two mathematical models and obtained the performance measures on the time to recruitment under different conditions. More specifically, in model 1 they have obtained these analytical results when the loss of manhours in the organization is the maximum of the loss of manhours for the two grades and the threshold for the organization is either the maximum or minimum of the thresholds for the loss of manhours in the two grades, assuming different distribution for these two thresholds for the grades. This work is extended in the present paper when the threshold for the loss of manhours in the organization is the sum of the corresponding thresholds of the two grades according as the two thresholds are exponential or extended exponential thresholds. This is done in section 2.1. In section 2.2, the performance measures are obtained when the threshold for the organization is minimum of the thresholds for the loss of manhours in the two grades where the loss of manhours in the two grades are put together to get the loss of manhours for the organization. The analytical results are numerically illustrated and the influence of nodal parameters on these results is studied and relevant conclusions are given.

## Model description and analysis for model -I

Consider an organization having two grades A and B , taking decisions at random epochs in $[0, \infty)$ and at every decision making epoch a random number of persons quit the organization. There is an associated loss of man power if the person quits. This loss of manhours linear and Cumulative and $X_{i}=\max \left(X_{i A}, X_{i B}\right)$. Let $X_{i}, i=1,2,3 \ldots \ldots$ be the loss of man hours due to the $\quad \mathrm{i}^{\text {th }}$ decision epoch in the organization and $\mathrm{G}($. [ $g($.$) ] be its distribution function[density function]. Let \mathrm{g}_{\mathrm{k}}($.$\left.) [ \mathrm{G}_{\mathrm{k}}().\right]$ be the k-fold convolution of $g().[G()$.$] . Let \bar{g}_{(.)}$be the Laplace transform of g (.). Let $\mathrm{S}_{\mathrm{k}}$ be the total loss of man power in first k decisions. It is assumed that the inter decision times are independent and identically distributed exponential random variables and $f().[F()$.$] be its density function [distribution$ function] with parameter $\theta$. Let $f_{k}().\left[F_{k}().\right]$ be the $k$-fold convolution of $f().[F()$.$] . Let V_{k}(t)$ be the probability that there are exactly ' $k$ ' decisions in $[0, t]$. Since the number of decision taken form a renewal process, from renewal theory it is known that $V_{k}(t)=F_{k}(t)-F_{k+1}(t)$ where $F_{0}(t)=1$. The loss of man power process and process of inter decision times are assumed to be statistically independent. Let $Y=Y_{A}+Y_{B}$ be the threshold level for the organization. For all $i=1,2,3 \ldots$, it is assumed that $X_{i}$ and $Y$ are independent. Recruitment is done whenever the cumulative loss of man hours in the organization crosses Y . Let W be the time for recruitment in the organization and $I$ (.) $[\mathrm{L}()$.$] be its density function[distribution function]. Let \mathrm{E}(\mathrm{W})$ and $\mathrm{V}(\mathrm{W})$ be the mean and variance of the time for recruitment.

## Main result

The survival function of $W$ is given by,
$P(W>t)=\sum_{k=0}^{\infty}\{$ Probability that there are exactly $k$ decisions in $[0, t)$ and cumulative loss of manhours does not crosses the threshold level $Y$ in these ' $k$ ' decisions \}

By the law of total probability

$$
\begin{aligned}
P(W>t)= & \sum_{k=0}^{\infty} V_{k}(t) P\left(S_{k}<Y=Y_{A}+Y_{B}\right) \\
= & \sum_{k=0}^{\infty} V_{k}(t) \\
& \int_{0}^{\infty} P\left(Y_{A}+Y_{B}>x\right) g_{k}(x) d x
\end{aligned}
$$

$=$
$\sum_{k=0}^{\infty}\left[F_{k}(t)-F_{k+1}(t)\right] \int_{0}^{\infty} P\left(Y_{A}+Y_{B}>x\right) g_{k}(x) d x$
Since $X_{i A}$ and $X_{B B}$ follow exponential distribution with parameters $\lambda_{A} \& \lambda_{B}$ respectively we find that $g(x)=$
$\lambda_{A} e^{-\lambda_{A} x}+\lambda_{B} e^{-\lambda_{B} x}-\left(\lambda_{A}+\lambda_{B}\right) e^{-\left(\lambda_{A}+\lambda_{B}\right) x}$.

## Case(i) :

Suppose $Y_{A}$ and $Y_{B}$ follow exponential distribution with parameters $\mu_{A \text { and }} \mu_{B}$ respectively. In this case it is found that

$$
\begin{align*}
& \int_{0}^{\infty} P\left(Y_{A}+Y_{B}>x\right) g_{k}(x) d x \frac{\mu_{A}}{\mu_{A}-\mu_{B}}\left[\mathrm{~A}_{1}\right]^{\mathrm{k}}- \\
& \frac{\mu_{B}}{\mu_{A}-\mu_{B}}\left[\mathrm{~A}_{2}\right]^{\mathrm{k}} \tag{3}
\end{align*} \ldots .
$$

where
$\mathrm{A}_{1}=\frac{\lambda_{A}}{\mu_{B}+\lambda_{A}}+\frac{\lambda_{B}}{\mu_{B}+\lambda_{A}}-\frac{\lambda_{A}}{\mu_{B}+\lambda_{A}+\lambda_{B}}$
$A_{2}=\frac{\lambda_{A}}{\mu_{A}+\lambda_{A}}+\frac{\lambda_{B}}{\mu_{A}+\lambda_{A}}-\frac{\lambda_{A}}{\mu_{A}+\lambda_{A}+\lambda_{B}}$

From (1) ,(3) and on simplification we get

$$
\begin{equation*}
\mathrm{L}(\mathrm{t})={ }_{1}-P(W>t)=\frac{\mu_{A}}{\mu_{A}-\mu_{B}} \mathrm{D}_{1}+\frac{\mu_{B}}{\mu_{A}-\mu_{B}} \mathrm{D}_{2} \tag{5}
\end{equation*}
$$

and

$$
\begin{equation*}
\bar{l}(s)=\frac{\mu_{A}}{\mu_{A}-\mu_{B}} \overline{a_{1}}+\frac{\mu_{B}}{\mu_{A}-\mu_{B}} \overline{a_{2}} \tag{6}
\end{equation*}
$$

where
for $\mathrm{j}=1,2, \mathrm{D}_{\mathrm{j}}=\mathrm{D}_{\mathrm{j}}(\mathrm{t})=\left(1-\mathrm{A}_{\mathrm{j}}\right) \sum_{k=0}^{\infty} F_{k}(t)\left[A_{j}\right]^{k-1}$ and

$$
\begin{equation*}
\bar{a}_{j}=\bar{a}_{j}(s)_{=\left(1-A_{\mathrm{j}}\right)} \sum_{k=0}^{\infty}\{\bar{f}(s)\}^{k}\left[A_{j}\right]^{k-1} . \tag{7}
\end{equation*}
$$

Since $\bar{f}(s)=\frac{\theta}{\theta+s}, \quad E(W)=-\left[\frac{d}{d s} \bar{l}(s)\right]_{s=0}$ and $E\left(W^{2}\right)=-\left[\frac{d^{2}}{d s^{2}} \bar{l}(s)\right]_{s=0}$
from (5),(6),(8) and on simplification we get

$$
\begin{equation*}
\mathrm{E}(\mathrm{~W})=\frac{1}{\theta}\left\{\frac{\mu_{A}}{\mu_{A}-\mu_{B}} C_{1}-\frac{\mu_{B}}{\mu_{A}-\mu_{B}} C_{2}\right\} \& \tag{9}
\end{equation*}
$$

$\mathrm{E}\left(\mathrm{W}^{2}\right)=\frac{2}{\theta^{2}}\left\{\frac{\mu_{A}}{\mu_{A}-\mu_{B}} C_{1}^{2}-\frac{\mu_{B}}{\mu_{A}-\mu_{B}} C_{2}^{2}\right\}$
where $\left.\mathrm{C}_{\mathrm{j}}=\frac{1}{\left[1-A_{j}\right]}\right], \mathrm{j}=1,2$, and $\mathrm{A}_{\mathrm{j}}$ 's are given by(4).
When $Y_{A}$ and $Y_{B}$ are exponential random variables (9) gives the mean time to recruitment. From (9) and (10) the variance of the time to recruitment can be computed for this case.

## Case(ii) :

Suppose $Y_{A}$ and $Y_{B}$ follow extended exponential distribution with scale parameters $\mu_{A}$ and $\mu_{B}$ respectively and shape parameter 2 . In this case it is found that

$$
\begin{array}{r}
\int_{0}^{\infty \infty} P\left(Y_{A}+Y_{B}>x\right) g_{k}(x) d x=\mathrm{B}_{1}[[1]]^{\mathrm{k}}-\mathrm{B}_{2}\left[[2]^{\mathrm{k}}\right. \\
+\mathrm{B}_{3}\left[[3]^{\mathrm{k}}-\mathrm{B}_{4}\left[[4]^{\mathrm{k}}\right.\right. \tag{11}
\end{array}
$$

## where

$$
\begin{align*}
& \mathrm{I}_{1}=\frac{\lambda_{A}}{\mu_{B}+\lambda_{A}}+\frac{\lambda_{B}}{\mu_{B}+\lambda_{A}}-\frac{\lambda_{A}}{\mu_{B}+\lambda_{A}+\lambda_{B}} \\
& \mathrm{I}_{2}=\frac{\lambda_{A}}{2 \mu_{A}+\lambda_{A}}+\frac{\lambda_{B}}{2 \mu_{A}+\lambda_{A}}-\frac{\lambda_{A}}{2 \mu_{A}+\lambda_{A}+\lambda_{B}} \\
& \mathrm{I}_{3}=\frac{\lambda_{A}}{\mu_{A}+\lambda_{A}}+\frac{\lambda_{B}}{\mu_{A}+\lambda_{A}}-\frac{\lambda_{A}}{\mu_{A}+\lambda_{A}+\lambda_{B}} \\
& \mathrm{I}_{4}=\frac{\lambda_{A}}{2 \mu_{B}+\lambda_{A}}+\frac{\lambda_{B}}{2 \mu_{B}+\lambda_{A}}-\frac{\lambda_{A}}{2 \mu_{B}+\lambda_{A}+\lambda_{B}} \tag{12}
\end{align*}
$$

$$
\mathrm{B}_{1}=\frac{4 \mu_{A}^{2}}{\left(2 \mu_{A}-\mu_{B}\right)\left(\mu_{A}-\mu_{B}\right)}
$$

$$
\begin{equation*}
\mathrm{B}_{2}=\frac{\mu_{B}^{2}}{\left(2 \mu_{A}-\mu_{B}\right)\left(\mu_{A}-\mu_{B}\right)} \tag{13}
\end{equation*}
$$

$$
\mathrm{B}_{3}=\frac{4 \mu_{B}^{2}}{\left(\mu_{A}-\mu_{B}\right)\left(\mu_{A}-2 \mu_{B}\right)}
$$

$$
\mathrm{B}_{1}=\frac{\mu_{A}^{2}}{\left(\mu_{A}-\mu_{B}\right)\left(\mu_{A}-2 \mu_{B}\right)}
$$

From (1), (11) and on simplification we get

$$
\mathrm{L}(\mathrm{t})=1-P(W>t)=\mathrm{B}_{1} \mathrm{D}_{1}-\mathrm{B}_{2} \mathrm{D}_{2}+\mathrm{B}_{3} \mathrm{D}_{3}-\mathrm{B}_{4} \mathrm{D}_{4} \ldots \text { (14) }
$$

and

$$
\bar{l}(s)=B_{2} \overline{a_{1}}-B_{2} \overline{a_{2}}+B_{3} \overline{a_{3}}-B_{4} \overline{a_{4}}
$$

Where, for $m=1,2,3,4$,

$$
\begin{align*}
& \mathrm{D}_{\mathrm{m}}=\mathrm{D}_{\mathrm{m}}(\mathrm{t})=\left(1-I_{\mathrm{m}}\right) \sum_{k=0}^{\infty} \quad F_{k}(t)\left[I_{m}\right]^{k-1} \\
& \text { and } \\
& \overline{a_{m}}=\overline{a_{m}}(s)=\left(1-I_{m}\right)\left[I_{m}\right]^{k-1} \tag{16}
\end{align*}
$$

Since $\bar{f}(s)=\frac{\theta}{\theta+s}, \quad \mathrm{E}(\mathrm{W})=-\left[\frac{d}{d s} \bar{l}(s)\right]_{s=0}$ and $E\left(\mathrm{~W}^{2}\right)=-\left[\frac{d^{2}}{d s^{2}} \bar{l}(s)\right]_{s=0}$
from (14),(15),(17) and on simplification we get

$$
\mathrm{E}(\mathrm{~W})=\frac{1}{\theta}\left\{B_{1} E_{1}-B_{2} E_{2}+B_{3} E_{3}-B_{4} E_{4}\right\}
$$

and

$$
\begin{equation*}
\mathrm{E}\left(\mathrm{~W}^{2}\right)=\frac{2}{\theta^{2}}\left\{B_{1} E_{1}^{2}-B_{2} E_{2}^{2}+B_{3} E_{3}^{2}-B_{4} E_{4}^{2}\right\} \tag{18}
\end{equation*}
$$

where $E_{m}=\frac{1}{\left[1-I_{m}\right]^{\prime}} m=1,2,3,4, I_{m}$ 's are given by(12) and $B_{m}$ 's are given by(13).

When $Y_{A}$ and $Y_{B}$ are extended exponential random variables (18) gives the mean time to recruitment. From (18) and (19) the variance of the time to recruitment can be computed for this case.

## Numerical illustration for model 1

The influence of nodal parameters on the performance measures namely mean and variance of the time to recruitment are studied numerically and the results are tabulated in Table-1. As per the Table-1, the researcher made the following observations.
(i) As $\lambda_{1}$ increases, the mean loss of man hours decreases and hence the mean of the time to recruitment increases. The variance of the time to recruitment also increases with $\lambda_{1}$.
(ii) As $\lambda_{2}$ increases, the mean loss of man hours decreases and hence the mean of the time to recruitment increases. The variance of the time to recruitment also increases with $\lambda_{2}$.
(iii) As $\theta$ increases, the mean inter-decision time decreases and hence the mean of the time to recruitment decreases. In this case the variance of the time to recruitment also decreases.

## Model description and analysis for model -2

All the assumptions and notations are same as in model 1 except the loss of man hours and the threshold for the organization. In this model we take the $X_{i}$ as the loss of manhours for the organization and $\mathrm{Y}=\min \left(\mathrm{Y}_{\mathrm{A}}, Y_{B}\right)$ as the threshold for the organization.

As in model-1, it can be shown that
By the law of total probability,
$P(W>t)=\sum_{k=0}^{\infty}\left[F_{k}(t)-F_{k+1}(t)\right]$
$\int_{0}^{\infty}[1-H(x)] g_{k}(x) d x$
Since $X_{i}$ follows exponential distribution with parameter $\lambda$ we find that $\mathrm{g}(\mathrm{x})=\lambda e^{-\lambda x}$

Suppose $Y_{A}$ and $Y_{B}$ follow extended exponential distribution with scale parameters $\mu_{A}$ and $\mu_{B}$ respectively and shape parameter 2 . In this case it is found that
$\int_{0}^{\infty}[1-H(x)] g_{k}(x) d x=\left[\mathrm{R}_{1}\right]^{\mathrm{k}}-2\left[\mathrm{R}_{2}\right]^{\mathrm{k}}-2\left[\mathrm{R}_{3}\right]^{\mathrm{k}}+4\left[\mathrm{R}_{4}\right]^{\mathrm{k}}$
where

$$
\begin{align*}
& \mathrm{R}_{1}=\bar{g}\left(2 \mu_{A}+2 \mu_{B}\right)=\frac{\lambda}{2 \mu_{A}+2 \mu_{B}+\lambda}  \tag{22}\\
& \mathrm{R}_{2}=\bar{g}\left(\mu_{A}+2 \mu_{B}\right)=\frac{\lambda}{\mu_{A}+2 \mu_{B}+\lambda}
\end{align*}
$$

$$
\begin{align*}
& \mathrm{R}_{3}=\bar{g}\left(2 \mu_{A}+\mu_{B}\right)=\frac{\lambda}{2 \mu_{A}+\mu_{B}+\lambda} \\
& \mathrm{R}_{4}=\bar{g}\left(\mu_{A}+\mu_{B}\right)=\frac{\lambda}{\mu_{A}+\mu_{B}+\lambda} \tag{23}
\end{align*}
$$

From (20) , (22) and on simplification we get
$\mathrm{L}(\mathrm{t})={ }_{1}-P(W>t)=\mathrm{D}_{1}-2 \mathrm{D}_{2}-2 \mathrm{D}_{3}+4 \mathrm{D}_{4}$
and
$\bar{l}(s)=\overline{a_{1}}-2 \overline{a_{2}}-2 \overline{a_{3}}+4 \overline{a_{4}}$
where for $m=1,2,3,4$
$\mathrm{D}_{\mathrm{m}}=\mathrm{D}_{\mathrm{m}}(\mathrm{t})=\left(1-\mathrm{R}_{\mathrm{m}}\right) \sum_{k=0}^{\infty} F_{k}(t)\left[R_{m}\right]^{k-1}$
and

$$
\begin{equation*}
\overline{a_{m}}=\overline{a_{m}}(s)=\left(1-R_{m}\right)\left[R_{m}\right]^{k-1} \tag{26}
\end{equation*}
$$

Since $\bar{f}(s)=\frac{\theta}{\theta+s}, \quad \mathrm{E}(\mathrm{W})=-\left[\frac{d}{d s} \bar{l}(s)\right]_{s=0}$ and
$\mathrm{E}\left(\mathrm{W}^{2}\right)=-\left[\frac{d^{2}}{d s^{2}} \bar{l}(s)\right]_{s=0}$
from (24),(25),(27) and on simplification we get
$\mathrm{E}(\mathrm{W})=\frac{1}{\theta}\left\{T_{1}-2 T_{2}-2 T_{3}+4 T_{4}\right\}$
and
$\mathrm{E}\left(\mathrm{W}^{2}\right)=\frac{2}{\theta^{2}}\left\{T_{1}^{2}-2 T_{2}^{2}-2 T_{3}^{2}+4 T_{4}^{2}\right\}$
where $T_{m}=\frac{1}{\left[1-R_{m}\right]^{\prime}}, m=1,2,3,4$. $R_{m}$ 's are given by(23).
When $Y_{A}$ and $Y_{B}$ are extended exponential random variables (28) gives the mean time to recruitment. From (28) and (29) the variance of the time to recruitment can be computed for this model.

## Numerical illustration for model-2

The influence of nodal parameters on the performance measures namely mean and variance of the three to recruitment are studied numerically and the results are tabulated in Table-2. As per the Table-2, the researcher made the following observations.
(i) As $\lambda$ increases, the mean loss of man hours decreases and hence the mean of the time to recruitment increases. The variance of the time to recruitment also increases with $\lambda$.
(ii) As $\theta$ increases, the mean inter-decision time decreases and hence the mean of the time to recruitment decreases. In this case the variance of the time to recruitment also decreases.

Table 1: Effect of $\boldsymbol{\lambda}_{1}, \boldsymbol{\lambda}_{2}$ and $\boldsymbol{\theta}_{\text {on performance measures }}$
$\left.\boldsymbol{\mu}_{\boldsymbol{1}=0.6,} \boldsymbol{\mu}_{\mathbf{2}}=0.4\right)$

| $\lambda_{1}$ | $\lambda_{2}$ | $\mu_{1}$ | $\mu_{2}$ | $\theta$ | case-1 |  | case-2 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | E(W) | V(W) | E(W) | V(W) |
| 0.1 | 0.2 | 0.6 | 0.4 | 0.10 | 12.2448 | 146.4712 | 13.7960 | 176.1273 |
| 0.3 | 0.2 | 0.6 | 0.4 | 0.10 | 14.8435 | 205.2293 | 17.9318 | 263.7550 |
| 0.5 | 0.2 | 0.6 | 0.4 | 0.10 | 16.1519 | 237.8067 | 19.8385 | 309.5899 |
| 0.3 | 0.2 | 0.6 | 0.4 | 0.10 | 14.8435 | 205.2293 | 17.9318 | 263.7550 |
| 0.3 | 0.4 | 0.6 | 0.4 | 0.10 | 17.5446 | 272.8019 | 22.1191 | 360.4827 |
| 0.3 | 0.6 | 0.6 | 0.4 | 0.10 | 19.1117 | 316.1597 | 24.4110 | 420.9452 |
| 0.3 | 0.2 | 0.6 | 0.4 | 0.10 | 14.8435 | 205.2293 | 17.9318 | 263.7550 |
| 0.3 | 0.2 | 0.6 | 0.4 | 0.15 | 9.8957 | 91.2130 | 11.9545 | 117.2245 |
| 0.3 | 0.2 | 0.6 | 0.4 | 0.20 | 7.4217 | 51.3073 | 8.9659 | 65.9388 |

Table 2: Effect of $\boldsymbol{\lambda}_{\text {and }} \boldsymbol{\theta}$ on performance measures

$$
\left(\boldsymbol{\mu}_{\boldsymbol{A}=0.6,} \boldsymbol{\mu}_{\boldsymbol{B}}=0.4\right)
$$

| $\lambda$ | $\theta$ | $\mu_{A}$ | $\mu_{B}$ | $\mathrm{E}(\mathrm{W})$ | $\mathrm{V}(\mathrm{W})$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.1 | 0.10 | 0.6 | 0.4 | 11.8214 | 138.0077 |
| 0.3 | 0.10 | 0.6 | 0.4 | 15.4643 | 223.4974 |
| 0.5 | 0.10 | 0.6 | 0.4 | 19.1071 | 321.6199 |
| 0.3 | 0.10 | 0.6 | 0.4 | 15.4643 | 223.4974 |
| 0.3 | 0.15 | 0.6 | 0.4 | 10.3095 | 99.3322 |
| 0.3 | 0.20 | 0.6 | 0.4 | 7.7321 | 55.8744 |

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