RRST-Statistics

# Forecasting Fish Product Export in Tamilnadu A Stochastic Model Approach 

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| Article Info |
| :--- |
| Article History |
| Received $:$ <br> Revised $\quad$ <br> Accepted $\quad$ |
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#### Abstract

This study aims at forecasting fish product export in Tamilnadu, based on data on inland and marine fish product export during the years from 1969 to 2008. The study considered Autoregressive (AR), Moving Average (MA) and Autoregressive Integrated Moving Average (ARIMA) processes to select the appropriate stochastic model for forecasting fish product export in Tamilnadu. Based on ARIMA (p, d, q) and its components ACF, PACF, Normalized BIC, Box-Ljung Q statistics and residuals estimated, ARIMA $(0,1,2)$ was selected. Based on the chosen model, it could be predicted that the fish product export would increase to 1 , 14,695 tons in 2015 from 74,549 tons in 2008 in Tamilnadu.


Key Words: Fish product export, BIC, Forecasting, ARIMA.

## Introduction

Fisheries play a vital role in providing protein-rich food at an affordable price to the people. The inland fisheries sector in Tamilnadu is spread over 3.71 lakh hectare of water-spread area comprising reservoirs, irrigation and seasonal tanks, ponds, estuaries and backwaters. Besides, the State has 56000 hectare of brackish water area suitable for aquaculture, of which, an area of 4455 hectare has already been developed for aquaculture. In Tamilnadu, maritime sector dominates the fishery sector, as the State has a coastal line of 1076 km ( $13.3 \%$ of the Nation's coast line of 8118 km ). The sector provides employment to 10.02 lakhs of marine and inland fisher-folk and contributes around Rs.20,000 crore to foreign exchange which accounts for $27.5 \%$ of exports from India. In this background, this study was conducted to forecast the future fish product export in the State, so as to help the policy planners to formulate needed strategies for achieving and sustaining the targets in the sector.

## Material and Methods

As the aim of the study was to forecast fish product export, various forecasting techniques were considered for use. ARIMA model, introduced by Box and Jenkins (1970) ${ }^{[1]}$, was frequently used for discovering the pattern and predicting the future values of the time series data. Akaike (1970) ${ }^{[2]}$ discussed the stationary time series by an $\operatorname{AR}(p)$, where $p$ is finite and bounded by the same integer. Moving Average (MA) models were used by Slutzky (1973)[3]. Hannan and Quinn (1979) ${ }^{[4]}$ suggested obtaining the order of a time series model by minimizing the errors for pure AR models, and Hannan (1980) ${ }^{[5]}$ for ARMA models. A second order determination method could be considered as a variance of Schwarz's Bayesian Criterion (SBC) which gives a consistent estimate of the order of an ARMA model. Hosking (1981) ${ }^{[6]}$ introduced a
family of models, called fractionally differenced autoregressive integrated moving average models, by generalizing the ' d ' fraction in ARIMA ( $p, \mathrm{~d}, \mathrm{q}$ ) model.

Stochastic time-series ARIMA models were widely used in time series data having the characteristics (Alan Pankratz, 1983[7]) of parsimonious, stationary, invertible, significant estimated coefficients and statistically independent and normally distributed residuals. When a time series is nonstationary, it can often be made stationary by taking first differences of the series i.e., creating a new time series of successive differences ( $\mathrm{Y}_{\mathrm{t}} \mathrm{Y}_{\mathrm{t}-1}$ ). If first differences do not convert the series to stationary form, then first differences can be created. This is called second-order differencing. A distinction is made between a second-order differences $\left(\mathrm{Y}_{\mathrm{t}}-\mathrm{Y}_{\mathrm{t}}\right.$ 2).

While Mendelssohn (1981) $)^{[8]}$ used Box-Jenkins ${ }^{[1]}$ models to forecast fishery dynamics, Prajneshu and Venugopalan (1996) ${ }^{[9]}$ discussed various statistical modeling techniques viz., polynomial, ARIMA time series methodology and nonlinear mechanistic growth modeling approach for describing marine, inland as well as total fish production in India during the period 1950-51 to 1994-95. Tsitsika et al. (2007) ${ }^{[10]}$ also used univariate and multivariate ARIMA models to model and forecast the monthly pelagic production of fish species in the Mediterranean Sea during 1990-2005. Jai Sankar et al. (2010) ${ }^{[11]}$ also used stochastic modeling for cattle production and forecast the yearly production of cattle in the Tamilnadu state during 1970-2010.

The time series when differenced follows both AR and MA models and is known as autoregressive integrated moving averages (ARIMA) model. Hence, ARIMA model was used in this study, which required a sufficiently large data set and
involved four steps: identification, estimation, diagnostic checking and forecasting. Model parameters were estimated using the Statistical Package for Social Sciences (SPSS) package and to fit the ARIMA models.

$$
\begin{aligned}
& \text { Autoregressive process of order } \quad \text { (p) is, } \\
& Y_{t}=\mu+\phi_{1} Y_{t-1}+\phi_{2} Y_{t-2}+\ldots .+\phi_{p} Y_{t-p}+\varepsilon_{t} ; \\
& \text { Moving Average process of order } \\
& Y_{t}=\mu-\theta_{1} \varepsilon_{t-1}-\theta_{2} \varepsilon_{t-2}-\ldots-\theta_{q} \varepsilon_{t-q}+\varepsilon_{t}
\end{aligned}
$$

and the general form of ARIMA model of order ( $\mathrm{p}, \mathrm{d}, \mathrm{q}$ ) is
$Y_{t}=\phi_{1} Y_{t-1}+\phi_{2} Y_{t-2}+\ldots+\phi_{p} Y_{t-p}+\mu-\theta_{1} \varepsilon_{t-1}-\theta_{2} \varepsilon_{t-2}-\ldots-\theta_{q} \varepsilon_{t-q}+\varepsilon_{t}$
where $Y_{\mathrm{t}}$ is fish product export, $\varepsilon_{t}$ 's are independently and normally distributed with zero mean and constant variance $\sigma^{2}$ for $t=1,2, \ldots, n$; $d$ is the fraction differenced while interpreting AR and MA and $\phi s$ and $\theta$ s are coefficients to be estimated.

Trend Fitting: The Box-Ljung Q statistics was used to transform the non-stationary data in to stationarity data and to check the adequacy for the residuals. For evaluating the adequacy of AR, MA and ARIMA processes, various reliability statistics like R2, Stationary R2, Root Mean Square Error (RMSE), Mean Absolute Percentage Error (MAPE), and Bayesian Information Criterion (BIC) were used. The reliability statistics viz. RMSE, MAPE, BIC and Q statistics were computed as below:

$$
\begin{aligned}
& R M S E=\left[\frac{1}{n} \sum_{i=1}^{n}\left(Y_{i}-\hat{Y}_{i}\right)^{2}\right]^{1 / 2} \\
& \text { MAPE }=\frac{1}{n} \sum_{i=1}^{n}\left|\frac{\left(Y_{i}-\hat{Y}_{i}\right)}{Y_{i}}\right|
\end{aligned}
$$

$\operatorname{BIC}(p, q)=\ln v^{*}(p, q)+(p+q)[\ln (n) / n]$
where $p$ and $q$ are the order of AR and MA processes
respectively and n is the number of observations in the time series and $\mathrm{v}^{*}$ is the estimate of white noise variance $\sigma^{2}$.

$$
Q=\frac{n(n+2) \sum_{i=1}^{k} r k^{2}}{(n-k)}
$$

where n is the number of residuals and rk is the residuals autocorrelation at lag k .

In this study, the data on fish product export in Tamilnadu were collected from the Department of Fisheries, Government of Tamilnadu for the period from 1969 to 2008 and were used to fit the ARIMA model to predict the future product export.

## Results and Discussion

Model Identification: ARIMA model was designed after assessing that transforming the variable under forecasting was a stationary series. The stationary series was the set of values that varied over time around a constant mean and constant variance. The most common method to check the stationarity is to explain the data through graph and hence is done in Figure 1.

Figure 1 reveals that the data used were non-stationary. Again, non-stationarity in mean was corrected through first differencing of the data. The newly constructed variable $Y_{t}$ could now be examined for stationarity. Since, $Y_{t}$ was stationary in mean, the next step was to identify the values of $p$ and q. For this, the autocorrelation and partial autocorrelation coefficients (ACF and PACF) of various orders of $Y_{t}$ were computed and presented in Table 1 and Figure 2.

The tentative ARIMA models are discussed with values differenced once ( $\mathrm{d}=1$ ) and the model which had the minimum normalized BIC was chosen. The various ARIMA models and the corresponding normalized BIC values are given in Table 2. The value of normalized BIC of the chosen ARIMA was 16.881.

Table 1. ACF and PACF of fish product export

| Lag | Auto <br> Correlation | Box-Ljung Statistic |  |  | Partial Auto <br> Correlation |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Value | Df | Sig. | Value | Df | Value | Df |
| $\mathbf{1}$ | -0.228 | 0.154 | 2.185 | 1 | 0.139 | -0.228 | 0.160 |
| $\mathbf{2}$ | 0.253 | 0.152 | 4.948 | 2 | 0.084 | 0.212 | 0.160 |
| $\mathbf{3}$ | -0.323 | 0.150 | 9.592 | 3 | 0.022 | -0.254 | 0.160 |
| $\mathbf{4}$ | 0.074 | 0.148 | 9.844 | 4 | 0.043 | -0.078 | 0.160 |
| $\mathbf{5}$ | 0.115 | 0.146 | 10.462 | 5 | 0.063 | 0.276 | 0.160 |
| $\mathbf{6}$ | 0.361 | 0.144 | 16.780 | 6 | 0.010 | 0.421 | 0.160 |
| $\mathbf{7}$ | -0.141 | 0.141 | 17.777 | 7 | 0.013 | -0.164 | 0.160 |
| $\mathbf{8}$ | 0.140 | 0.139 | 18.791 | 8 | 0.016 | 0.036 | 0.160 |
| $\mathbf{9}$ | -0.471 | 0.137 | 30.622 | 9 | 0.000 | -0.268 | 0.160 |
| $\mathbf{1 0}$ | 0.259 | 0.135 | 34.326 | 10 | 0.000 | 0.018 | 0.160 |
| $\mathbf{1 1}$ | -0.052 | 0.132 | 34.483 | 11 | 0.000 | 0.063 | 0.160 |
| $\mathbf{1 2}$ | 0.178 | 0.130 | 36.353 | 12 | 0.000 | -0.151 | 0.160 |

Table 2. BIC values of ARIMA (p, d, q)

| ARIMA $(\mathbf{p}, \mathrm{d}, \mathrm{q})$ | BIC values |
| :--- | :--- |
| $(0,1,0)$ | 16.945 |
| $(0,1,1)$ | 16.940 |
| $(0,1,2)$ | 16.881 |
| $(1,1,0)$ | 16.955 |
| $(1,1,1)$ | 17.012 |
| $(1,1,2)$ | 17.113 |
| $(2,1,0)$ | 17.069 |
| $(2,1,1)$ | 17.19 |
| $(2,1,2)$ | 17.064 |



Figure 1. Time plot of fish product export in Tamilnadu


Figure 2. ACF and PACF of differenced data


Figure 3. Residuals of ACF and PACF


Figure 4. Actual and estimate of fish product export
Model Estimation: Model parameters were estimated using SPSS package and the results of estimation are presented in Tables 3 and 4. $R^{2}$ value was 0.97 . Hence, the most suitable model for fish production was ARIMA (0, 1, 2), as this model had the lowest normalized BIC value, good $R^{2}$ and better model fit statics (RMSE and MAPE).

Table 3. Estimated ARIMA model of fish production

|  | Estimate | SE | $\mathbf{t}$ | Sig. |
| :--- | :--- | :--- | :--- | :--- |
| Constant | -260934.103 | 24933.056 | -10.465 | 0.000 |
| MA 1 | 0.739 | 162.592 | 0.005 | 0.996 |
| MA 2 | 0.261 | 42.491 | 0.006 | 0.995 |

Table 4. Estimated ARIMA model fit statistics

| Fit Statistic | Mean |
| :--- | :--- |
| Stationary R-squared | 0.318 |
| R-squared | 0.976 |
| RMSE | 3837.419 |
| MAPE | 17.219 |
| Normalized BIC | 16.881 |

Table 5. Residual of ACF and PACF of fish product export

| Lag | ACF |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Mean | SE | PACF |  |
| Lag 1 | 0.049 | 0.160 | Mean | SE |
| Lag 2 | 0.168 | 0.161 | 0.049 | 0.160 |
| Lag 3 | -0.386 | 0.165 | 0.166 | 0.160 |
| Lag | -0.112 | 0.187 | -0.413 | 0.160 |
| Lag 5 | -0.059 | 0.188 | -0.099 | 0.160 |
| Lag 6 | 0.214 | 0.189 | 0.123 | 0.160 |
| Lag 7 | -0.179 | 0.195 | 0.120 | 0.160 |
| Lag 8 | -0.031 | 0.199 | -0.385 | 0.160 |
| Lag 9 | -0.517 | 0.199 | -0.086 | 0.160 |
| Lag 10 | 0.063 | 0.231 | -0.358 | 0.160 |
| Lag 11 | -0.062 | 0.232 | 0.029 | 0.160 |
| Lag 12 | 0.224 | 0.232 | -0.051 | 0.160 |


| Year | Actual | Predicted | LCL | UCL |
| :---: | :---: | :---: | :---: | :---: |
| 1969 | 5220 | -- | -- | -- |
| 1970 | 5637 | 4601 | -5079 | 14282 |
| 1971 | 5942 | 4800 | -4310 | 13910 |
| 1972 | 6765 | 4913 | -3742 | 13569 |
| 1973 | 5511 | 5384 | -3056 | 13824 |
| 1974 | 2572 | 4976 | -3316 | 13269 |
| 1975 | 5719 | 3989 | -4202 | 12181 |
| 1976 | 3800 | 5378 | -2739 | 13495 |
| 1977 | 4720 | 4694 | -3365 | 12754 |
| 1978 | 5743 | 5505 | -2509 | 13519 |
| 1979 | 6945 | 6155 | -1822 | 14131 |
| 1980 | 8294 | 7077 | -869 | 15023 |
| 1981 | 8550 | 8139 | 219 | 16059 |
| 1982 | 7000 | 8952 | 1054 | 16850 |
| 1983 | 6872 | 9307 | 1428 | 17186 |
| 1984 | 6252 | 10219 | 2357 | 18082 |
| 1985 | 18792 | 10900 | 3052 | 18748 |
| 1986 | 18456 | 15880 | 8045 | 23715 |
| 1987 | 18053 | 16379 | 8555 | 24202 |
| 1988 | 16745 | 18026 | 10213 | 25839 |
| 1989 | 15330 | 19107 | 11303 | 26910 |
| 1990 | 22768 | 20287 | 12492 | 28082 |
| 1991 | 26851 | 24141 | 16353 | 31928 |
| 1992 | 24949 | 26634 | 18853 | 34414 |
| 1993 | 30963 | 27869 | 20095 | 35642 |
| 1994 | 20311 | 31772 | 24004 | 39540 |
| 1995 | 28831 | 30257 | 22495 | 38019 |
| 1996 | 31330 | 35527 | 27770 | 43284 |
| 1997 | 40878 | 37592 | 29840 | 45344 |
| 1998 | 41052 | 42699 | 34951 | 50446 |
| 1999 | 45026 | 44600 | 36857 | 52344 |
| 2000 | 43464 | 48486 | 40746 | 56225 |
| 2001 | 53005 | 50391 | 42655 | 58127 |
| 2002 | 58482 | 56031 | 48298 | 63763 |
| 2003 | 70147 | 59822 | 52093 | 67551 |
| 2004 | 68462 | 66061 | 58335 | 73787 |
| 2005 | 70809 | 68141 | 60418 | 75864 |
| 2006 | 72418 | 72438 | 64717 | 80158 |
| 2007 | 72883 | 76025 | 68307 | 83743 |
| 2008 | 74549 | 79533 | 71818 | 87248 |


| 2009 | - | 83443 | 75730 | 91156 |
| :--- | :--- | :--- | :--- | :--- |
| 2010 | - | 89378 | 81360 | 97397 |
| 2011 | - | 94178 | 86159 | 102196 |
| 2012 | - | 99109 | 91091 | 107127 |
| 2013 | - | 104172 | 112190 |  |
| 2014 | - | 109368 | 10155 | 117385 |
| 2015 | - | 114695 | 122712 |  |

Diagnostic Checking: The model verification is concerned with checking the residuals of the model to see if they contained any systematic pattern which still could be removed to improve the chosen ARIMA, which has been done through examining the autocorrelations and partial autocorrelations of the residuals of various orders. For this purpose, various autocorrelations up to 12 lags were computed and the same along with their significance tested by Box-Ljung statistic are provided in Table 5. As the results indicate, none of these autocorrelations was significantly different from zero at any reasonable level. This proved that the selected ARIMA model was an appropriate model for forecasting fish product export in Tamilnadu.

The ACF and PACF of the residuals are given in Figure 3, which also indicated the 'good fit' of the model. Hence, the fitted ARIMA model for the fish product export data was:

$$
Y_{t}=-260934.103-0.739 \varepsilon_{t-1}-0.261 \varepsilon_{t-2}+\varepsilon_{t}
$$

Forecasting: Based on the model fitted, forecasted fish product export (in tons) for the year 2009 through 2015 respectively were 83443, 89378, 94178, 99109, 104172, 109368 and 114695 tons (Table 6). To assess the forecasting ability of the fitted ARIMA model, the measures of the sample period forecasts' accuracy were also computed. This measure also indicated that the forecasting inaccuracy was low. Figure 4 shows the actual and forecasted value of fish product export (with $95 \%$ confidence limit) in the State.

## Conclusion

The most appropriate ARIMA model for fish product export forecasting was found to be ARIMA (0, 1, 2). From the forecast available from the fitted ARIMA model, it can be found that forecasted product export would increase to $1,14,695$ tons in 2015 from 74,549 tons in 2008. That is, using time series data from 1969 to 2008 on fish product export, this study provides evidence on future fish product export in the State,
which can be considered for future policy making and formulating strategies for augmenting and sustaining fish product export in the State.

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